1	Economic optimization of wildfire intervention activities
2	
3	
4	David T. Butry, Jeffrey P. Prestemon, Karen L. Abt, Ronda Sutphen ¹

¹ David T. Butry is Economist, Office of Applied Economics, Building and Fire Research Laboratory, National Institute of Standards and Technology, 100 Bureau Drive, Mail Stop 8603, Gaithersburg, MD 20899-8603—tel. 301-975-6136, <u>david.butry@nist.gov</u> (*corresponding author*); .Jeffrey P. Prestemon is Research Forester, Southern Research Station, Forestry Sciences Laboratory, PO Box 12254, Research Triangle Park, NC, 27709—tel. 919-549-4033, jprestemon@fs.fed.us; Karen L. Abt is Research Economist, Southern Research Station, Forestry Sciences Laboratory, PO Box 12254, Research Triangle Park, NC, 27709—tel. 919-549-4094, <u>kabt@fs.fed.us;</u> Ronda Sutphen is Cooperative Fire Protection Manager, Forest Protection, Florida Division of Forestry, Department of Agriculture and Consumer Services, 3125 Conner Boulevard, Tallahassee, Florida 32399-1650—tel. 850-488-6111, <u>sutpher@doacs.state.fl.us</u>.

Economic optimization of wildfire intervention activities

5 6

-

7 Abstract

8 We describe how two important tools of wildfire management, wildfire prevention education and 9 prescribed fire for fuels management, can be coordinated to minimize the combination of 10 management costs and expected societal losses resulting from wildland fire. We present a long-11 run model that accounts for the dynamics of wildfire, the effects of fuels management on wildfire 12 ignition risk and area burned, and the effects of wildfire prevention education on the ignition risk 13 of human-caused, unintentional wildfires. Based on wildfire management activities in Florida 14 from 2002 to 2007, we find that while wildfire prevention education and prescribed fire have 15 different effects on timing and types of fires, the optimal solution is to increase both interventions. Prescribed fire affects whole landscapes and therefore reduces losses from all 16 17 wildfire types (including lightning), while wildfire prevention education reduces only human-18 caused ignitions. However, prescribed fire offers a longer-term solution with little short-term 19 flexibility. Wildfire prevention education programs, by comparison, are more flexible, both in 20 time and space, and can respond to unexpected outbreaks, but with limited mitigation longevity. 21 Only when used together in a coordinated effort do we find the costs and losses from 22 unintentional wildfires are minimized. 23

25 I. Introduction

Wildfires are produced on a landscape from a combination of purchased and free inputs. Free (i.e., non-market) inputs to wildfire include lightning ignitions and those caused by humans,

27 (i.e., non market) inputs to whethe mende ingitting ignitions and those edused by numans,

28 natural fuels (vegetation), and weather conditions. Purchased inputs include anything employed

by fire managers to affect fire occurrence, extent, and intensity. Wildfire managers operate in a

30 world of constraints to their actions to affect wildfire processes, so the decisions made are

31 typically choices among competing means of intervening in wildfire processes.

32

33 Economic theory (e.g., Rideout and Omi 1990) provides a framework for understanding the

34 effects of decisions and quantifying the trade-offs among alternative actions: under risk

35 neutrality, minimizing the sum of management costs incurred and the expected losses

36 experienced by society from wildfires that occur. In economics, at the optimum, the cost of the

37 last unit of each purchased input reduces the expected losses by the identical amount. Because

inputs and wildfires themselves have both short- and long-run impacts on costs and losses, this
 economic expression of optimality—and hence purchased input trade-offs—is inherently long-

40 run (e.g., Mercer et al. 2007).

41

42 A challenge in empirical wildfire economics is obtaining the information needed to quantify the 43 marginal contributions among alternative fire management actions, enabling better decision-44 making. This article describes how two important purchased inputs of wildfire management, 45 wildfire prevention education and prescribed fire, can be used in combination to achieve the 46 economic objective of minimizing long-run management costs and expected societal losses. We 47 describe a long-run model that accounts for the dynamics of wildfire, which provides fuel 48 reduction as a free input in subsequent fire seasons; the short- and long-run effects of fuels 49 management on fire extent and occurrence; and the short- and long-run effects of wildfire 50 prevention education on the occurrence on targeted unintentional wildfires.

51

52 This paper makes the following contributions to the literature. First, we outline a model that 53 incorporates both fire ignitions and prescribed fire in an economic model of wildfire 54 management. Second, we describe how wildfire prevention and prescribed fire trade off in the 55 pursuit of an optimal policy. Prescribed fire operates over whole landscapes and therefore affects 56 the losses associated with all fire types, while fire prevention only operates directly on part of 57 potential fire starts. Previous research has focused on individually optimizing either fuels 58 management activities (e.g., Kim et al. 2009; Wei et al. 2008; Mercer et al. 2007; Yoder 2004) or 59 suppression resources (Haight et al. 2007; Donovan 2006; Donovan and Rideout 2003; 60 MacLellan and Martell 1996), so to minimize the expected losses of wildfire. Joint 61 optimizations have been explored, but these have focused on optimizing between a 62 preoperational and an operational phase (Minciardi et al. 2009). An example would be to jointly 63 optimize fuels management and suppression effort (e.g., Drucker et al. 2008, Mercer et al. 2008). 64 We, instead, optimize over two preoperational wildfire management strategies while holding 65 suppression effort constant. Third, we show that the quantities of free inputs affect trade-offs 66 and optimal amounts of purchased inputs in wildfire management, implying that the optimal 67 combinations of purchased inputs should vary, along with the variation in free inputs, both over 68 time and across space.

69

The organization of the rest of the manuscript is as follows: Section II presents our theoretical model of wildfire management economics; Section III describes the study site and the two wildfire management variables of interest (wildfire prevention education and prescribed fire treatments); Section IV introduces the empirical model of wildfire ignition risk and Section V describes the optimization methodology; Sections VI and VII present the empirical and optimization results; and Section VIII provides the conclusion.

76

77 II. A Theoretical Model

The expected cost plus loss of wildfire is the sum of expected ignitions times expected fire size times the loss value per acre, and the sum of all the intervention costs. Let $I_{i,t}^{p}$ be the count of ignitions of targeted unintentional fire types (i.e., human-caused, unintentional ignitions targeted by prevention education activities) in location i (i = 1 to J) in period t (t= 1 to T); $I_{i,t}^{n}$ be the count of other ignitions (i.e., other unintentional, intentional, naturally-occurring wildfire ignitions) in iand t; \mathbf{x}_{i}^{p} be a vector of an unspecified number of lags of wildfire prevention actions in period t; \mathbf{x}_{i}^{R} be a vector of an unspecified number of lags of other actions (e.g., prescribed fire); \mathbf{z}_{i} be an unspecified number of lags of free inputs to wildfire production in period t. Thus, targeted

86 unintentional and other ignitions can be represented as

87

88
$$I_{i,t}^{p} = f(\mathbf{x}_{i}^{p}, \mathbf{x}_{i}^{R}, \mathbf{z}_{i})$$

- 89 And
- 90 $I_{i,t}^n = f(\mathbf{x}_i^R, \mathbf{z}_i)$
- 91

92 The size of wildfires, $A_{i,t}$, is a function of lagged values of prescribed fire and free inputs, as 93 prevention inputs do not influence fire size, and can be represented by

94

95
$$A_{i,t} = A(\mathbf{x}_i^R, \mathbf{z}_i)$$

96

97 Let w^p be an index of the cost of wildfire prevention actions, $x_{i,t}^p$ the quantity of those actions in 98 period *t*, w^R be an index of the price of other actions, $x_{i,t}^R$ the quantity of those other actions in 99 period *t*, so that the costs of intervention are

100

101
$$C_{i,t} = w^p x_{i,t}^p + w^R x_{i,t}^R$$

102

103 $Y_{i,t}^{p}$ and $Y_{i,t}^{n}$ are the loss per acre of targeted unintentional and non-targeted wildfire. In addition, 104 *E* is the expectations operator and *r* is the discount rate. Thus, the fire management problem is: 105

106
$$\min_{x_i^p, x_i^r} M = \sum_{i}^{J} \sum_{t}^{T} (1+r)^{-t} \{ C_{i,t} + Y_{i,t}^p E[I_{i,t}^p A_{i,t}^p] + Y_{i,t}^n E[I_{i,t}^n A_{i,t}^n] \}$$
(1)

107

108 where *M* is the expected cost plus loss of wildfire. As written, fire prevention efforts affect 109 only $I_{i,t}^{p}$ while the other inputs to the fire production process (prescribed fire and free) affect all 110 ignitions as well as the expected fire sizes of both types of fires.

- 112 The optimal allocation of wildfire prevention education, $(x_{i,t}^{p^*})$ across space and time and the
- analogous allocation of prescribed fire $(x_{i,t}^{R^*})$ would yield a long-run minimum of the objective
- 114 function (minimizing cost plus loss) at Z^* . At the optimum, the partial derivative of Z^* with
- 115 respect to $x_{i,t}^{p^*}$ should equal the unit price of those efforts, or $\partial Z^* / \partial x_{i,t}^{p^*} = w^p$; similarly,
- 116 $\partial Z^* / \partial x_{i,t}^{R^*} = w^p$. Depending on the specification of the ignition process, free inputs may affect
- 117 optimal levels of purchased inputs (i.e., a non-linear in parameters functional form). For
- 118 example, a Poisson specification of the ignition process implies that inputs are non-separable and
- 119 thus optimal input quantities are jointly determined.
- 120

121 III. Wildfire Interventions

Wildfire prevention education (WPE), defined here as the avoidance of targeted unintentional human caused wildfires through education,² includes activities such as radio, television, and newspaper public service announcements (PSAs); homes visitations; presentations; fliers and brochures distributed; and community wildland-urban interface hazard assessments. We also explored the effect of prescribed fire fuel treatments, those specifically targeted towards reducing wildfire hazards, on targeted unintentional ignitions. WPE and prescribed fire offer land managers different mechanisms to minimize the impact of future wildfire.

129

130 We explored the effect of these two interventions across the four wildfire management regions in 131 Florida (Figure 1). Region 1 includes 16 counties in the panhandle of Florida, as well as the 132 cities of Tallahassee and Pensacola and, along with Region 2, represents the primary timber 133 growing region of the state. The 18 counties in Region 2 are home to both the city of 134 Jacksonville and the extreme southern part of the Okefenokee Swamp. Region 3 includes 15 135 counties in central Florida, including the cities of Orlando, Daytona, and Tampa. The 136 southernmost region, Region 4, includes Lake Okeechobee, the Everglades, the city of Miami 137 and the Keys in its 18 counties.

² These include debris fire escapes, campfire escapes, and fires caused by discarded cigarettes and by children. We ignore other kinds of unintentional fire starts (such as equipment and railroad fires) because they are not the focus of wildfire prevention education, and we ignore arson because its occurrence is affected by a different combination of managerial (and law enforcement) actions (e.g., Prestemon and Butry 2005).

140 Over the study period (2002 to 2007), Florida experienced 6338 targeted unintentional ignitions 141 accounting for 96 830 acres burned. The number of targeted unintentional wildfire ignitions 142 varied between 20/month in Region 3 to 37/month in Region 2. The number of acres burned 143 varied between 255 acres/month in Region 1 to 828 acres/month in Region 2. Region 2 144 experienced more than twice the amount of burned acres than the next fire-prone region (Region 145 4). While targeted unintentional ignitions made up 37 % of all wildfire ignition reported over 146 this period, targeted unintentional wildfires remained small. They comprised only 7 % of the 147 total burned acres. Historically, targeted unintentional wildfires have accounted for more acres 148 burned (natural fires may burn larger areas due to changes in climate and weather). However, 149 because targeted unintentional ignitions are caused by humans, these wildfires tend to occur in 150 places close to values at risk (e.g., Bradshaw 1988, Butry et al. 2002). 151 152 Over the study period more than 1.4 million acres burned from wildfire. Another 2.8 million

acres were authorized for burning by silvicultural-based prescribed fire treatments targeting
hazardous fuels. The number of prescribed fire permits issued varied from as low as 28/month in
Region 4 to as high as 149/month in Region 1, on average. Region 1 also averaged the most
requested number acres for treatment, at 23 016 acres/month, compared to Region 2 with 6487
acres/month. On average, monthly prescribed fire fuel treatments involve 8 to 90 times more
acres than do wildfires.

159

160 The intensity and mix of WPE activities vary by wildfire management region (see Table 2).

161 Distributing wildfire prevention fliers, brochures, and CDs appeared to be the most common

activity across regions (176 452 were distributed in all). Public service announcements (PSAs)

163 were also very common (30 931). Overall, television PSAs (12 504) were most widely used

164 followed by newspaper (11 020) and radio (7407) spots. Also used were 7314 homes visitations,

165 890 presentations, and 156 community WUI hazard assessments.

166

167 Timing is important when developing (pre-fire) mitigation strategies. Figure 2 presents the

168 average seasonality of targeted unintentional and non-targeted unintentional (e.g., arson and

lightning) wildfire ignitions, authorized prescribed fire acres and WPE activities over the 2002 to

170 2007 study period. Shown is the monthly count of each data series compared to its 12-month 171 average value. Targeted unintentional wildfire ignitions peaked in the late winter and early 172 spring, as did authorized prescribed fire treatments, fire prevention brochure distribution, and 173 wildfire education presentations. Media PSAs and home visitations peaked prominently in late 174 spring and early summer. Hazard assessments did not follow any strong seasonal trend. 175 Interestingly, media PSAs and home visitations peaked after the peak of targeted unintentional 176 ignitions. In fact, the relatively large peaks exhibited by media PSAs and home visitations were 177 followed with declines in targeted unintentional ignitions, perhaps indicative of their 178 effectiveness at reducing wildfire risks. Non-targeted unintentional ignitions peaked during this 179 period, providing an indication that climatological and fuel conditions in the summer improve 180 wildfire ignition success. Likely this explains why prescribed fire authorizations also were fewer 181 during this fire-prone period (i.e., higher likelihood of escaped prescribed fires). 182 183 Casually, it appears wildfire mitigation effort reduced targeted unintentional ignitions, as periods 184 of high effort were followed by periods of lower targeted unintentional ignitions. Of course, it

185 also looks as if high periods of effort were accompanied by high periods of ignitions, so likely 186 there is some simultaneous determination occurring. Our statistical model, presented in the next 187 section, untangled the complicated relationships between wildfire and prevention by accounting 188 for endogeneity and other factors related to the ignition generation process (e.g., weather, fire 189 history, and socioeconomic characteristics of the spatial units of inference).

190

191 IV. Empirical Model

192 The statistical model estimates the effect of free inputs (including the weather, vegetation and 193 climate) and purchased inputs (WPE and prescribed fire) on the monthly occurrence of targeted 194 unintentional wildfires across the four fire management regions. We assume the occurrence of 195 reported targeted unintentional wildfire follows a Poisson distribution:

196

197
$$I_{i,t}^{p} = e^{\mathbf{a}_{i}^{\prime} \mathbf{z}_{i,t} + \mathbf{\beta}_{i}^{\prime} \mathbf{x}_{i,t-k} + \varepsilon_{i,t}}$$
(2)

where $I_{i,t}^{p}$ is the number of targeted unintentional wildfires for location *i* in time *t*, **z** are the free 199 200 inputs to wildfire production, **x** are the *M* interventions occurring over the current and *k* previous 201 months, α and β are the parameters associated with the inputs and interventions, respectively, 202 and ε is an error term. Because of simultaneity between the number of targeted unintentional 203 wildfires and interventions, the inputs to wildfire production are correlated with the error term, 204 $E[\varepsilon_{m,i,t} | x_{m,i,t}] \neq 0$. Thus, we augment equation (2) with a set of auxiliary equations, called "control functions" to construct a set of variables to control for the unobserved heterogeneity 205 206 creating bias in (2) (see Hausman 1978):

207

$$208 x_{m,i,t} = \mathbf{\gamma}'_{m,i,t} \mathbf{h}_{m,i,t} + c_{m,i,t} (3)$$

209

where **h** is a set of instruments and **c** is a normally-distributed error term. Procedurally, the
controls are obtained by regressing intervention effort on the set of instruments and estimating
the residuals, so that,

213

214
$$\hat{c}_{m,i,t} = x_{m,i,t} - \hat{\gamma}'_{m,i,t} \mathbf{h}_{m,i,t}$$
 (4)

215

- 216 Equation (2) is augmented to become:
- 217

218
$$I_{i,t}^{p} = e^{\boldsymbol{a}_{i}^{\prime}\boldsymbol{z}_{i,t} + \boldsymbol{\beta}_{i}^{\prime}\boldsymbol{x}_{i,t-k} + \boldsymbol{\delta}_{i}^{\prime}\hat{\boldsymbol{e}}_{i,t} + \boldsymbol{\xi}_{i,t}}$$
(5)

219

where ξ is a normally-distributed error term and by construction it is not correlated with the inputs to wildfire production (i.e., $E[\xi_{m,i,t} | x_{m,i,t}, \hat{c}_{m,i,t}] = 0$). We used maximum likelihood estimation to maximize the log-likelihood function based on (5) :

223

224
$$\ln L = \sum_{i=1}^{I} \sum_{t=1}^{T} -e^{\boldsymbol{\alpha}'_{i} \mathbf{z}_{i,t} + \boldsymbol{\beta}'_{i} \mathbf{x}_{i,t-k} + \boldsymbol{\delta}'_{i} \hat{\mathbf{c}}_{i,t} + \boldsymbol{\xi}_{i,t}} + \left(\boldsymbol{\alpha}'_{i} \mathbf{z}_{i,t} + \boldsymbol{\beta}'_{i} \mathbf{x}_{i,t-k} + \boldsymbol{\delta}'_{i} \mathbf{c}_{i,t}\right) I_{i,t}^{p} - \ln(I_{i,t}^{p}!)$$
(6)

226 The intervention variables, $\mathbf{x}_{i,m,i,t-k}$, include WPE variables for current and k=6 lagged months (a 227 vector that includes the individual sums of the WPE variables over the previous 6 months) and 228 the area of prescribed fire permits issued in the previous one, two, and three years. The WPE 229 variables include the number of media public service announcements (TV, radio, and print ads) 230 (Media), homes visited (Homes), presentations given (Presentations), brochures and flyers 231 distributed (Brochures), and community wildfire hazard assessments (Hazard) provided in 232 current month t and over the last six months (Florida Division of Forestry 2008a). Although 233 several other WPE measures (fairs, billboards, movie theater public service announcements) 234 were undertaken by wildfire prevention specialists, the occurrence of such measures was too 235 sparse to allow for identification. All included WPE variables were normalized by population, 236 but population was included as an additional explanatory variable in the statistical models to 237 account for the changes in the levels of the integer Poisson process. The other intervention 238 variables include the annual area authorized for hazard removal (as opposed to for ecological or 239 wildlife reasons) by prescribed burning lagged up to three years (Florida Division of Forestry 240 2008b).

241

242 The vector of free inputs, $\mathbf{z}_{i,i,t}$, includes measures of fire weather (relative humidity (RH, current 243 month and 12 month lag), Keetch-Byram Drought Index (KBDI, current month and 12 month 244 lag; Keetch and Byram [1968]), Fire Weather Index (FWI, current month and 12 month lag; 245 Fosberg [1978]), Modified Fire Weather Index (MFWI, current month and 12 month lag; 246 Goodrick [2002]), precipitation) (Goodrick 2008), climate (the March to September monthly 247 average and the October to February monthly average of the Niño-3 sea surface temperature 248 anomaly in degrees centigrade, [National Oceanic and Atmospheric Administration 2008]), the 249 annual area burned (in acres) by wildfire lagged up to six years (Florida Division of Forestry 250 2008c), county population estimates (U.S. Bureau of the Census 2008), the number of sworn 251 full-time equivalent police officers per capita (Florida Department of Law Enforcement 2008), 252 and dummy variables for region (Region 1 is included in the intercept), season (fall is included in 253 the intercept), and year (2002 is included in the intercept). Finally, we include a trend variable to 254 account for the net effects of unspecified steady changes not captured by other variables. 255

256 The vector of instruments included all of the variables used in the prevention models except 257 current WPE activities (in this model the dependent variable), and also included wildfire 258 ignitions of targeted unintentional causes (lagged 2 years to 5 years) and the one-year lagged 259 value of sales tax revenues (Sales Tax) (Florida Department of Revenue 2008). These variables 260 were chosen as instruments based on our assumption that they are correlated with WPEs but not 261 with current wildfire behavior, except through their effect on WPE. For instance, prior wildfire 262 behavior could influence future WPE strategies, and sales tax revenues could influence future 263 WPE by affecting WPE budgets.

264

265 V. Optimal Mitigation

We assumed that a prevented fire reduced the number of fires in the same location and the same
month and year of the average size as the fires that occurred in that month and location.
Interventions affect wildfire acres burned through two methods: (1) the effect of prevention on
targeted unintentional ignitions (current model); and (2) the effect of prescribed fire on area
burned for fires that occur (Mercer et al. 2007 model).

271

We simulated the effects of changes in prevention efforts and prescribed fire (*X*) on targeted unintentional ignitions (I^{p}) and area burned (*A*). In the long-run the change in area burned (A^{*}) equals the sum of the change in the long-run area burned ignited by non-targeted sources (A^{*n}) and the change in the long-run area burned ignited by targeted unintentional sources (A^{*n}) is the sum of the long-run area burned ignited by targeted unintentional sources

277

$$278 \qquad \Delta A^* = \Delta A^{*n} + \Delta A^{*p} \tag{7}$$

279

This is ultimately a function of the short-run change in area of targeted unintentional wildfire due to prevention change (A_t^p) :

282

283
$$\Delta A^* = -0.633 * \Delta A_t^p + (1 - 0.633) * \Delta A_t^p + \sum_k \frac{\Delta A_t^p}{\Delta A_{t-k}}$$
(8)

where we accounted for the fuel treatment effect of wildfire. An average acre in period *t* increases wildfire in the long-run by about 0.633 (Mercer et al. 2007). The short-run change in area of targeted unintentional wildfire due to a prevention change (ΔA_t^p) is:

288

$$\Delta A_t^p = \frac{\partial I_t^p}{\partial X} * \overline{A}_{i,t}$$
(9)

290

where $\partial I_t^p / \partial X$ is determined via estimation of (5) $(\partial I_t^p / \partial X = \beta)$ and $\overline{A}_{i,t}$ is the average size of the targeted unintentional fires that occurred in the same month, year, and fire management region.

294

We explored three scenarios: (1) minimize cost plus loss by altering WPE, holding prescribed fire constant; (2) minimize cost plus loss by altering prescribed fires, holding WPE constant; and (3) minimize cost plus loss by altering both WPE and prescribed fires. Losses from wildfire were set at \$1267/acre burned (per Mercer et al. [2007] adjusted to 2005 dollars).³

299

300 Florida's annual wildfire prevention education budget is \$0.47 million. The annual budget

301 allocation across wildfire management regions is not known with precision; however, it is

302 believed the allocation is roughly equivalent across regions (R. Rhea, Florida Division of

303 Forestry, pers. comm., October 24, 2008). We explored the sensitivity of this assumption by

304 examining the change when the spending was allocated proportionally based on historical

305 targeted unintentional wildfire acres burned.

306

The annual cost of prescribed fire fuel treatments is about \$3.2 million/year and they are largely borne by both private landowners and government. We assume a unit price of \$25/acre (based on an approximation from Cleaves et al. [2000]), but also explore the effect a price sensitive unit cost has on the results (i.e., prices rises with demand). Mercer et al. (2007) found that the

³ This figure assumes a constant cost plus loss per acre of wildfire. An alternative assumption, allowing costs plus losses to have a fixed cost per fire and a variable cost per acre burned, was not testable with the available data.

elasticity of the prescribed fire service supply with respect to price was 0.54 in Florida, and thatthe short-run wildfire area elasticity with respect to prescribed fire area was -0.73.

313

314 VI. Statistical Results

315 The control function models are significant and the covariates explain as much as 25 % to 52 % 316 of the variation in the WPE variables (see Table 3). The constructed control function variables 317 were used as additional model regressors in the targeted unintentional wildfire ignition model. 318 They have significant positive correlations (at the 10 % level) with targeted unintentional 319 ignitions, meaning endogeneity exists between WPE and targeted unintentional ignition rates 320 (see Table 4). The positive correlations imply that a standard Poisson regression estimation 321 would produce biased downward treatment effects on the WPE variables. The ignition model is significant and based on the calculated pseudo-R², explains 72 % of the variation in targeted 322

323 unintentional ignition counts.

324

325 Media PSAs, presentations, brochures, and community hazard assessments are significant at the 326 (10 % level) and negatively related to targeted unintentional wildfire ignition occurring in the 327 same month, after accounting for endogeneity. Home visitations are only weakly correlated 328 (13 % level). Lagged levels (activity within the last six months) of media PSAs, presentations, 329 and brochures are also significant (10 % level) and negatively related to ignitions. The 330 implication is that media PSAs, presentations, and brochures have both immediate and short-331 term mitigation effects, whereas community hazard assessments have an immediate effect, but 332 no lasting impact. Authorized prescribed fire acres have longer term effects, as compared to 333 WPE. Prescribed fire had a beneficial statistical effect (10 % level) two and three years post-334 treatment; however, prescribed fire performed within the last year did not have an impact on 335 targeted unintentional ignitions. (This does not rule out treatment effects on other types of 336 ignitions; this was not explored.) Other estimated relationships produced expected signs and 337 significance. Weather, climate, seasonality, historical fire patterns, and socioeconomic variables 338 are correlated with targeted unintentional ignitions, as are differences across regions and years. 339

The elasticities associated with prescribed fire fuel treatments (two and three year lagged) are
larger, in absolute terms, than most of the WPE variables, although in many cases the difference

342 is small. The elasticity associated with media PSAs (normalized by population) ran over the last

343 six months (-0.26) is the same as the elasticity associated with prescribed fire treatments

344 performed two years prior (-0.26). Thus, a 20 % increase in PSAs and prescribed fire would

have each decreased ignitions by 5.2 %, or on average 1.5 ignitions. This 20 % increase would

have required either an additional 118 PSAs or 2140 acres treated. The nonlinearity of the

- 347 Poisson model also assumes that WPE and fuel treatments are interdependent, thus the amount of
- 348 fuel treatment applied impacts the effect WPE had on ignition success (and vice-versa).
- 349

350 VII. Optimal Mitigation Results

351 Optimal Change in WPE Spending (Only)

352 The optimal change in WPE spending, holding prescribed fire constant, is a 225 % increase,

353 statewide (Figure 3). This figure shows that large increases in WPE would be needed in all four

regions to minimize costs plus losses under the two assumptions of initial equal or initial

- 355 proportional spending allocation.
- 356

357 The relative large increases when moving from the assumption of equal allocation to 358 proportional, in Regions 1 (+74 %) and 3 (+81 %), is likely due to the assumed initial small 359 spending level with a proportional allocation compared to the equal allocation. Under the 360 proportionality assumption Regions 1 and 3 initially receive 15 % and 16 %, respectively. 361 Region 2 receives 49 %, Region 3 receives 16 %, and Region 4 receives 20 % of the initial 362 spending allocation. Thus, Regions 1 and 3 appear to produce the greatest return on WPE 363 investment, and hence the substantial need for increased funding. The return on WPE also looks 364 more favorable for Region 4 under the proportionality assumption. Expansion of WPE in 365 Regions 1, 3, and 4 come at the expense of Region 2, which begins with a high initial allocation 366 level under the proportionality assumption, and quickly experiences larger diminishing returns. 367

368 Optimal Change in Prescribed Fire (Only)

369 The optimal change in prescribed fire, holding WPE spending constant, is a 79 % increase,

370 statewide (Figure 4). Results are nearly identical regardless of the prescribed fire unit cost price

- assumption. Optimality results in a 17 % decrease for Region 1, a 28 % increase for Region 3, a
- 372 122 % increase for Region 2, and a 180 % increase for Region 4. On average, Region 1

373 performed substantially more prescribed fire treatments (23 016 acres/month) over the observed 374 study period than any of the other regions—nearly 2.5 times the amount of the next largest 375 region (see Table 1). While on average Region 4 treated the second most acres (9134) and 376 performed the most number of WPE activities (individually and as a whole) per month, it also 377 experienced far more wildfire (by any cause). Prescribed fire affects wildfire regardless of 378 ignition. So, this explains the substantial increase in prescribed fire in the region. Over the study 379 period Region 4 experienced an average fire size of 151.5 acres; Region 2 was second with an 380 average size of 26.9 acres, followed by Region 3 (average equal to 18.2 acres) and Region 1 381 (average equal to 14.9 acres). Looking at the historical annual number of acres burned, this 382 ordering is preserved: Region 4—152 891 acres/year; Region 2—57 201 acres/year; Region 3— 383 22 870 acres/year; and Region 1—13 984 acres/year. With less wildfire, from all causes, 384 Regions 1 and 3 have less need to increase prescribed fire.

385

386 Optimal Change in Wildfire Interventions (Both)

Previously we explored the optimal change in one prevention strategy while holding the other fixed. Those solutions are useful when one strategy can be varied (i.e., additional funding) while the other faces the status quo. The optimal solution will result when both strategies (prescribed fire and WPE) can adjust. As we show below, it does not always lead to an expansion of both strategies. Given the functional form of ignition processes and the feedbacks that wildfires have on aggregate fuels levels, the optimal levels of both sets of inputs (WPE and prescribed fire) are determined jointly.

394

395 The optimal change in WPE and prescribed fire, assuming equal allocation of initial WPE

396 spending and price responsive prescribed fire service, is a 168 % increase in WPE and 74 %

increase in prescribed fire, statewide (Figure 5.) Region 1 faces the most extreme changes: a

398 304 % increase in WPE and a 29 % decrease in prescribed fire. Region 3 faces a similar pattern:

a 251 % increase in WPE and a 22 % increase in prescribed fire. Regions 2 and 4 fall in

400 between. Both require roughly a doubling of WPE and prescribed fire effort.

401

402 The optimal overall statewide change in WPE and prescribed fire, assuming proportional403 allocation of initial WPE spending and price responsive prescribed fire service, is identical to the

solution for an equal allocation (Figure 6). The relative increases are similar to that with the
equal allocation assumption. In fact, the initial allocation assumption does not affect the optimal
level of prescribed fire. When the proportionality assumption is maintained, WPE expenditures
are expanded over the case with an equal allocation assumption for Regions 1, 2 and 4. These
expansions come at the expense of Region 2 where the optimal increase is reduced from 162 %
to 136 %.

410

411 Tradeoff Analysis

412 Comparing the joint optimization strategy to the single strategy (holding the other input fixed) 413 shows that the optimal increases in statewide WPE and prescribed fire are less than that required 414 when one of the inputs is held fixed. In the joint optimization, the optimal increase in WPE is 415 168 % and 74 % for prescribed fire. Compare this to the case when prescribed fire is held fixed, 416 the optimal increase in WPE is 225 %, or when WPE is held fixed, the optimal increase in 417 prescribed fire is 79 %. In all regions, except Region 2, joint optimization results in a smaller 418 relative increase in WPE and prescribed fire. Joint optimization results in a larger relative 419 increase in WPE for Region 2, however. (Region 2 experienced the most unintentional wildfires 420 in terms of ignitions and acres burned; see Table 1).

421

422 In the joint optimization, there is a tradeoff between WPE and prescribed fire. While WPE is 423 effective, it targets only targeted unintentional ignitions; whereas prescribed fire targets all 424 wildfire types, regardless of the ignition source. This indiscriminate targeting mitigates the loss 425 of the "fuel treatment effect" of wildfire caused by ignition prevention because prescribed fire 426 still impacts the burn area of those wildfires that do occur. The optimal level of pre-fire 427 mitigation occurs through joint optimization of WPE and prescribed fire, requiring an increase of 428 WPE by 168 % and 74 % for prescribed fire. Joint optimization is preferred to single 429 optimization as it produces an expected cost plus loss lower than any produced through single 430 estimation (Table 5). Based on a statewide allocation strategy (i.e., increasing WPE and 431 prescribed fire equally across regions), the expected cost plus loss is \$301 million, a savings of 432 \$24 million (Table 5; Figure 7). Based on a regional allocation strategy (i.e., varying the 433 increase of WPE and prescribed fire across regions), the expected cost plus loss is further

reduced to \$287 million, a savings of \$38 million (Table 5). These savings are net saving and
already account for (offset) increased program costs.

436

437 VIII. Conclusion

438 We examined the effect of WPE and prescribed fire, two alternative pre-fire intervention 439 strategies, on targeted unintentional ignitions in Florida from 2002 to 2007. These targeted 440 unintentional ignitions included those occurring from escaped debris fires, escaped campfire, 441 fires caused by discarded cigarettes, and by children. During the study period, targeted 442 unintentional ignitions accounted for 37 % of all wildfire ignitions, but only 7 % of acres burned. Leveraging the measured effect of WPE and prescribed fire on targeted unintentional ignitions 443 444 and on the observed sizes of wildfires based on previous studies, we simulated changes in the 445 intervention levels to identify their optimal levels and the corresponding expected cost plus loss 446 due to wildfire damage. Expected cost plus loss was minimized with an increase in WPE of 447 168 % and prescribed fire acres treated of 74 %.

448

449 While these levels may be optimal, they may not be feasible. In fact, the State may not have the 450 ability to dramatically alter the scale of prescribed fire programs, unlike WPE, in Florida due to 451 land ownership limitations. Only a portion of at-risk forests are under State (or other 452 governmental) control, and these would where prescribed fire could most easily expanded by government policy.⁴ Constraints on prescribed fire, related to weather or smoke, may also limit 453 454 its expansion to levels less than 74 %. Relatedly, prescribed fires usually occur early in the 455 calendar year, and while our results suggest benefits last for several years, they also require a 456 year to take effect (at least statistically). At-risk areas must be identified well ahead of the threat. 457 The effect of WPE that we found in our modeling is shorter lived than prescribed fire, as we only 458 found a six-month maximum effect; however, there is evidence that WPE could be used 459 successfully to respond to outbreaks of targeted unintentional ignitions. Media PSAs, 460 presentations, brochures and fliers, and WUI community hazard assessments were found to 461 reduce the number of targeted unintentional ignitions in the same month that they were 462 performed. A 10 % increase in WPE was shown to have a 1.2 % to 2.3 % decrease in targeted

⁴ A program focusing on private lands would require a prescribed fire incentive program, which we did not evaluate in this study.

463 unintentional ignitions of the same month. Longer-term (up to six months) effects were shown 464 to occur for media PSAs, presentations, and brochures and fliers. In addition to the 1.4 % to 465 2.3 % real-time decrease in targeted unintentional ignitions from a 10 % increase in these education strategies, another 2.2 % to 2.6 % decrease in targeted unintentional ignitions would 466 467 be expected over the next six months. A 10 % increase in media PSAs, for example, is expected 468 to have 4.9 % reduction in targeted unintentional ignitions over a seven month period. This 469 marginal effect is on the order of magnitude of prescribed fire. Prescribed fire offers a longer-470 term solution at the expense of short-term flexibility. On the other hand, wildfire prevention 471 education programs offer the flexibility, both in time and space, to respond to outbreaks. When 472 used together in a coordinated effort, the program costs and wildfire damages from targeted 473 unintentionally set fires are minimized.

474

475 Refinements of our analyses could be pursued. We chose a simple analysis that asked how much 476 greater or lesser effort should be expended to minimize the sum of costs and expected losses 477 from wildfire in Florida. But a time-varying optimization analysis could also have been explored: 478 how much should WPE or prescribed fire efforts been changed over each of the units of time of 479 our analysis to achieve a minimum of costs plus losses? Further, we chose to change all WPE 480 activities simultaneously, assuming that absolute levels of each may vary only together, not 481 independently. However, given that each WPE type has a different observed effect on targeted 482 unintentional ignitions, it makes sense that the fire manager would prefer to allocate efforts 483 across types to achieve optimal fire management outcomes. Finally, our analysis was backward-484 looking. A forward-looking analysis might simulate future quantities of free inputs and identify 485 optimal stationary quantities of WPE and prescribed fire that would achieve minimum long-run 486 discounted costs plus losses, along the lines of Mercer et al. (2007). Given that absolute amounts 487 of free inputs vary across space in Florida, that analysis would identify differential amounts and 488 paths of future expected fire across fire regions in the state.

- 489 **References**
- 490 Bradshaw, W.G. 1988. Fire protection in the wildland/urban interface: Who plays what role?
 491 Fire Technology 24(3):195-203.
- 492
- 493 Butry, D.T., J.M. Pye, and J.P. Prestemon. 2002. Prescribed fire in the interface: Separating the
- 494 people from the trees. P. 132-136 in Outcalt, K.W. (ed.) 2002. Proceedings of the eleventh
- biennial southern silvicultural research conference. US For. Serv. Gen. Tech. Rep. SRS-48. 611
- 496 р.
- 497
- 498 Cleaves, D.A., J. Martinez, and T.K. Haines. 2000. Influences on Prescribed Burning Activity
- 499 and Costs in the National Forest System. USDA For. Serv. Gen. Tech. Rpt. SRS-37. Asheville,
- 500 NC: U.S. Department of Agriculture, Forest Service, Southern Research Station. 34 p.
- 501

502 Donovan, G.H. 2006. Determining the optimal mix of federal and contract fire crews: a case 503 study from the Pacific Northwest. Ecological Modeling 194(4):372-378.

- 504
- 505 Donovan, G.H., and D.B. Rideout. 2003. An integer programming model to optimize resource 506 allocation for wildfire containment. Forest Science 49(2):331-335.
- 507
- 508 Drucker, A.G., S.T. Garnett, M.K. Luckert, G.M. Crowley, and N. Gobius. 2008. Manager-based
- 509 valuation of alternative fire management regimes on Cape York Peninsula, Australia.
- 510 International Journal of Wildland Fire 17(5):660-673.
- 511
- 512 Florida Department of Law Enforcement. 2008. Sworn police officer data, 1989-2007. Data
- 513 obtained by special request, February 14, 2008.
- 514
- 515 Florida Division of Forestry. 2008a. Fire prevention activities by wildfire mitigation specialist by
- 516 month, paper and electronic records, 1999-2007. Data obtained April 23, 2008.
- 517
- 518 Florida Division of Forestry. 2008b. Prescribed fire permits issued, electronic records, 1989-
- 519 2007. Data obtained August 22, 2008.

520	
521	Florida Division of Forestry. 2008c. Wildfire activity, electronic records, 1980-2007. Data
522	obtained June 13, 2008.
523	
524	Florida Department of Revenue. 2008. Tax collections from July 2003. Available at
525	http://dor.myflorida.com/dor/taxes/colls_from_7_2003.html; last accessed December 5, 2008.
526	
527	Fosberg, M.A. 1978. Weather in wildland fire management: the fire weather index. P. 1-4 in
528	Proc. Conference on Sierra Nevada Meteorology, Lake Tahoe, CA.
529	
530	Goodrick, S.L. 2002. Modification of the Fosberg fire weather index to include drought. Int. J.
531	Wildl. Fire 11:205-221.
532	
533	Goodrick, S.L. 2008. Data provided by special request.
534	
535	Haight, R.G., and J.S. Fried. 2007. Deploying wildland fire suppression resources with a
536	scenario-based standard response model. INFOR: Informational Systems and Operational
537	Research 45(1):31-39.
538	
539	Hausman, J.A. 1978. Specification tests in econometrics. <i>Econometrica</i> 46(6):1251-1271.
540	
541	Keetch, J.J., and G.M. Byram. 1968. A drought index for forest fire control. US For. Serv. Res.
542	Pap. SE-38. 32 p.
543	
544	Kim, Y.H., P. Bettinger, and M. Finney. 2009. Spatial optimization of the pattern of fuel
545	management activities and subsequent effects on simulated wildfires. European Journal of
546	Operational Research 197(1):253-265.
547	
548	MacLellan, J.I., and D.L. Martell. 1996. Basing airtankers for forest fire control in Ontario.
549	Operations Research 44(5):677-686.
550	

551	Mercer, D.E., R.G. Haight, and J.P. Prestemon. 2008. Analyzing trade-offs between fuels
552	management, suppression, and damages from wildfire. P. 247-272 In Holmes, T.P., J.P.
553	Prestemon, and K.L. Abt (eds.), The Economics of Forest Disturbances: Wildfires, Storms, and
554	Invasive Species. Springer: Dordrecht, The Netherlands.
555	
556	Mercer, D.E., J.P. Prestemon, D.T. Butry, and J.M. Pye. 2007. Evaluating alternative prescribed
557	burning policies to reduce net economic damages from wildfire. Amer. J. Agric. Econ. 89(1):63-
558	77.
559	
560	Minciardi, R., R. Sacile, and E. Trasforini. 2009. Resource allocation in integrated preoperational
561	and operational management of natural hazards. Risk Analysis 29(1):62-75.
562	
563	National Oceanic and Atmospheric Administration. 2008. El Niño-Southern Oscillation
564	sea surface temperatures. Available online at
565	ftp.cpc.ncep.noaa.gov/wd52dg/data/indices/sstoi.indices; accessed Dec. 8, 2008.
566	
567	Prestemon, J.P., and D.T. Butry. 2005. Time to burn: modeling wildland arson as an
568	autoregressive crime function. American Journal of Agricultural Economics 87(3):756-770.
569	
570	Rideout, D.B., and P.N. Omi. 1990. Alternate expressions for the economic theory of forest fire
571	management. For. Sci. 36(3):614-624.
572	
573	US Bureau of the Census. 2008. Population estimates. Available at
574	http://www.census.gov/popest/counties/; last accessed June 2, 2008.
575	
576	US Department of Commerce. 2008. Consumer price index for all urban consumers, not
577	seasonally adjusted, monthly. Available at http://146.142.4.24/; last accessed February 5, 2008.
578	
579	Wei, Y., D. Rideout, A. Kirsch. 2008. An optimization model for locating fuel treatments across
580	a landscape to reduce expected fire losses. Canadian Journal of Forest Research 38(4):868-877.

- 582 Yoder, J. 2004. Playing with fire: endogenous risk in resource management. American Journal of
- 583 Agricultural Economics 86(4):933-948.

586 **Tables and Figures**

- 587 Table 1. Monthly number of targeted unintentionally-ignited wildfires and acres burned, and
- 588 prescribed fire (for hazard reduction) permits issued and acres treated in Florida from 2002 to
- 589 2007, by regions.

	Average	Minimum	Maximum	Observations
Region 1				
Targeted Unintentional Wildfire Ignitions	27	1	128	58
Targeted Unintentional Wildfire Acres	255	0.1	2490	58
Prescribed Fire Permits	149	2	836	58
Prescribed Fire Acres	23,016	3	127,878	58
Region 2				
Targeted Unintentional Wildfire Ignitions	37	2	139	57
Targeted Unintentional Wildfire Acres	828	0.4	35,639	57
Prescribed Fire Permits	85	1	420	57
Prescribed Fire Acres	6487	3	32,508	57
Region 3				
Targeted Unintentional Wildfire Ignitions	20	1	78	60
Targeted Unintentional Wildfire Acres	260	0.1	3989	60
Prescribed Fire Permits	41	0	160	60
Prescribed Fire Acres	9050	0	35,578	60
Region 4				
Targeted Unintentional Wildfire Ignitions	26	0	97	57
Targeted Unintentional Wildfire Acres	337	0	3156	57
Prescribed Fire Permits	28	0	98	57
Prescribed Fire Acres	9134	0	37,660	57

590

592 Table 2. Monthly wildfire prevention education activities recorded by wildfire mitigation

	Average	Minimum	Maximum	Observations
Region 1				
Radio PSAs	44	2	143	58
TV PSAs	5	0	48	58
Newspaper PSAs	5	0	39	58
Homes visited	96	0	1923	58
Presentations offered	0.3	0	1	58
Fliers, brochures, or CDs	162	0	1935	58
WUI hazard assessments	0.2	0	1	58
Region 2				
Radio PSAs	38	0	704	57
TV PSAs	59	0	911	57
Newspaper PSAs	75	0	1181	57
Homes visited	9	0	210	57
Presentations offered	2	0	23	57
Fliers, brochures, or CDs	904	0	3400	57
WUI hazard assessments	0.8	0	13	57
Region 3				
Radio PSAs	7	0	42	60
TV PSAs	23	0	147	60
Newspaper PSAs	14	0	83	60
Homes visited	4	0	115	60
Presentations offered	6	0	37	60
Fliers, brochures, or CDs	275	0	1897	60
WUI hazard assessments	0.6	0	6	60
Region 4				
Radio PSAs	41	0	283	57
TV PSAs	131	10	1630	57
Newspaper PSAs	99	0	2031	57
Homes visited	16	0	500	57
Presentations offered	6	0	109	57
Fliers, brochures, or CDs	1737	0	24500	57
WUI hazard assessments	1	0	9	57

593 specialists in Florida 2002 to 2007, by regions.

	MEDIA		HOMES		PRESENT	BROCHURES		HAZARI
	Coeff.		Coeff.		Coeff.	Coeff.		Coeff.
Ignitions: 2 year lag	1.32E-07		1.61E-07	*	-9.21E-10	-8.83E-08		-8.97E-1
Ignitions: 3 year lag	1.17E-07		7.37E-08		-2.96E-10	-3.44E-08		-4.72E-10
Ignitions: 4 year lag	1.66E-07	**	9.74E-09		1.17E-09	-2.68E-07		-6.44E-10
Ignitions: 5 year lag	1.11E-07	*	-6.41E-08		9.17E-10	-3.10E-07		-4.26E-1
MFWI: 12 month lag	-5.44E-06		-2.73E-06		-1.31E-07	-1.24E-05		-4.98E-0
FWI: 12 month lag	8.84E-07		6.62E-07		1.22E-07	-2.14E-05		3.61E-0
RH: 12 month lag	1.66E-06		-3.26E-06	*	-5.38E-09	-5.32E-06		-1.40E-0
KBDI: 12 month lag	1.03E-07		1.05E-07		8.37E-10	2.32E-07		1.81E-1
Sales Tax: 1 year lag	2.88E-14		-5.09E-15		5.64E-16	-6.91E-14		-3.56E-1
FWI :current?	-2.24E-05	*	2.03E-05		3.34E-07	2.42E-05		1.35E-0
RH	2.07E-06		8.47E-07		5.11E-08	3.23E-06		8.66E-0
KBDI	-1.57E-07		2.81E-07	**	2.36E-09	7.18E-07		1.04E-0
MFWI	3.11E-05	**	-1.51E-05		-2.99E-07	-4.54E-06		-1.52E-0
Niño 3: March-September	1.82E-06		-1.39E-05		1.69E-07	-3.52E-05		-1.30E-0
Niño 3: October-February	-7.80E-06		-5.20E-06		2.05E-07	4.43E-05		3.02E-0
Precipitation	-1.64E-06		-2.05E-07		-6.06E-08	-7.84E-06		-1.00E-0
Rx Fire: 1 year lag	6.37E-11		3.65E-10		-9.02E-12	2.00E-09		-2.94E-1
Rx Fire: 2 year lag	-1.34E-10		8.17E-10	***	-3.50E-12	2.18E-10		-2.72E-1
Rx Fire: 3 year lag	1.57E-10		-1.50E-10		-4.47E-12	-1.56E-09		-1.10E-1
Fire: 1 year lag	1.04E-10		-1.07E-10		2.92E-12	8.76E-11		9.05E-1
Fire: 2 year lag	-7.23E-10	*	-1.01E-09	**	6.83E-12	-2.86E-09	*	-1.44E-1
Fire: 3 year lag	-4.43E-10	*	-3.52E-10		3.73E-12	-2.95E-11		3.90E-1
Fire: 4 year lag	-3.24E-10	*	-1.39E-10		1.22E-12	1.51E-09	*	-6.89E-1
Fire: 5 year lag	-3.25E-10		-4.37E-10	*	3.12E-12	-1.20E-10		-1.49E-1
Fire: 6 year lag	-2.20E-10		-3.69E-10	**	2.03E-12	-2.87E-10		4.25E-1
Region 2	5.50E-03	**	-1.52E-03		-1.34E-05	1.69E-03		6.15E-0
Region 3	7.70E-03	***	3.13E-04		-1.31E-05	2.98E-03		1.09E-0
Region 4	8.45E-03	***	1.26E-03		-1.53E-05	3.97E-03		1.29E-0
Spring	5.57E-05	**	-3.24E-05		6.99E-07	7.77E-05		1.02E-0
Summer	2.40E-05		4.60E-05		4.73E-07	5.40E-05		2.11E-0
Winter	-2.10E-06		-1.04E-05		4.00E-07	1.19E-04		-1.61E-0

Table 3. Control function equation estimates for five prevention education variables.

Population	-5.69E-10	**	-3.90E-10		-5.92E-13	-3.42E-10		-9.10E-13	
Police per capita	1.75E+00	**	-6.56E-01		-3.84E-03	7.56E-02		1.59E-03	
2003	-3.19E-06		3.08E-05		-9.35E-07	-4.11E-06		-1.53E-07	
2004	3.65E-05		5.48E-05		-1.49E-06	3.19E-04		-5.47E-08	
2005	3.99E-05		7.80E-05		-1.72E-06	1.01E-03	***	-3.13E-07	
2006	6.75E-05		8.69E-05		-6.14E-07	1.02E-03	**	-1.55E-07	
2007	7.56E-05		1.70E-04		-3.12E-07	1.30E-03	**	3.72E-07	
Trend	1.25E-05	**	-1.79E-06		7.08E-08	-3.39E-05		6.58E-09	
Media: 1-6 months prior	-2.55E-01	***	-8.91E-03		-6.60E-04	5.13E-01	**	-7.61E-04	*
Homes: 1-6 months prior	-2.06E-01		4.52E-01	***	6.18E-04	1.09E+00	*	4.75E-04	
Presentations: 1-6 months prior	-4.15E+00	*	-3.61E+00		-4.21E-02	2.59E+00		-5.77E-04	
Brochures: 1-6 months prior	-1.74E-03		2.02E-03		3.42E-05	-1.39E-01	***	-9.76E-05	*
Hazard: 1-6 months prior	4.13E+01	***	-1.05E+01		-4.44E-02	2.55E+01		-5.42E-02	
Intercept	-8.76E-03	**	3.43E-03		1.62E-05	6.19E-04		-3.35E-06	
Prob > F	0.0000		0.0000		0.0671	0.0000		0.0000	
R^2	0.5155		0.3742		0.2471	0.4367		0.3973	

595 ***, **, * Denotes significances at the 0.01, 0.05, 0.10 levels, respectively.

Table 4. Poisson model estimate of the count of targeted unintentional wildfires, 2002 to 2007,

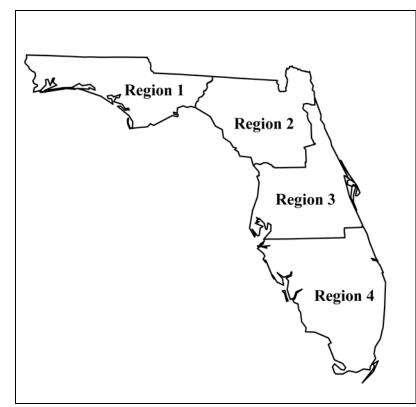
	Coeff.	Std. Error	Z	P> z	Elasticity
FWI	1.46E-01	6.11E-02	2.39	0.017	1.06
RH	-3.32E-02	8.72E-03	-3.8	0	-1.70
KBDI	1.66E-03	5.72E-04	2.9	0.004	0.41
MFWI	-5.25E-02	5.96E-02	-0.88	0.379	-0.32
Nino 3: March	2.98E-02	0.058193	0.51	0.609	-0.01
Nino 3: October	4.41E-02	0.0548857	0.8	0.422	0.02
Precipitation	-1.21E-01	1.33E-02	-9.13	0	-0.55
Rx Fire: 1 year lag	-1.41E-06	1.40E-06	-1.01	0.311	-0.18
Rx Fire: 2 year lag	-2.44E-06	1.44E-06	-1.7	0.089	-0.26
Rx Fire: 3 year lag	-3.66E-06	9.96E-07	-3.67	0	-0.34
Fire: 1 year lag	2.57E-06	7.79E-07	3.3	0.001	0.12
Fire: 2 year lag	-7.03E-06	1.36E-06	-5.15	0	-0.35
Fire: 3 year lag	-2.17E-06	8.37E-07	-2.59	0.01	-0.14
Fire: 4 year lag	1.42E-07	8.99E-07	0.16	0.875	0.01
Fire: 5 year lag	-1.88E-06	6.72E-07	-2.8	0.005	-0.21
Fire: 6 year lag	-1.60E-06	4.34E-07	-3.69	0	-0.18
Region 2	2.85E+01	1.03E+01	2.77	0.006	7.01
Region 3	4.62E+01	1.42E+01	3.25	0.001	11.94
Region 4	5.34E+01	1.58E+01	3.37	0.001	13.11
Spring	9.24E-01	1.37E-01	6.76	0	0.24
Summer	6.59E-01	1.18E-01	5.6	0	0.15
Winter	5.09E-01	1.10E-01	4.61	0	0.13
Population	-4.53E-06	1.20E-06	-3.78	0	-19.95
Police per capita	8.12E+03	3.21E+03	2.53	0.011	22.17
2003	6.71E-01	1.56E-01	4.29	0	0.14
2004	2.18E+00	1.94E-01	11.25	0	0.45
2005	3.31E+00	3.37E-01	9.81	0	0.68
2006	4.81E+00	3.73E-01	12.87	0	0.99
2007	6.26E+00	4.53E-01	13.84	0	0.65
Trend	-4.28E-02	1.88E-02	-2.28	0.023	-1.26
Media: 1-6 months prior	-1.34E+03	5.87E+02	-2.29	0.022	-0.26

597 and associated elasticities, calculated at the mean of the data.

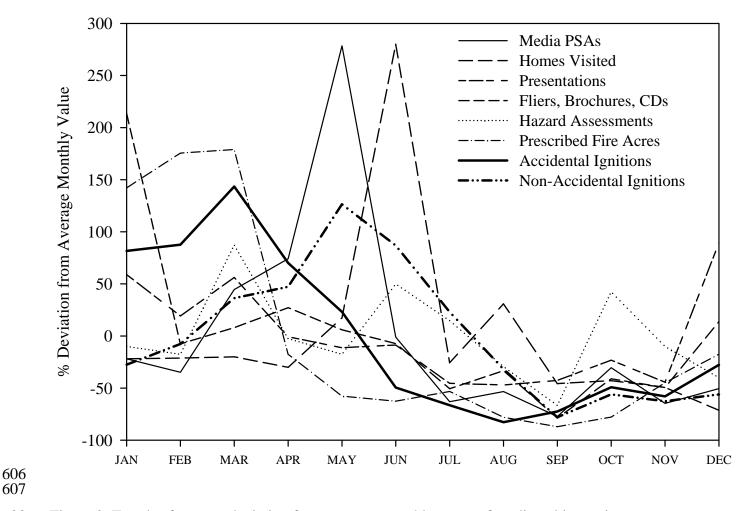
Homes: 1-6 months prior	4.50E+02	7.02E+02	0.64	0.521	0.04
Presentations: 1-6 months prior	-4.89E+04	8.69E+03	-5.63	0	-0.22
Brochures: 1-6 months prior	-2.15E+02	4.66E+01	-4.62	0	-0.24
Hazard: 1-6 months prior	6.54E+04	5.03E+04	1.3	0.194	0.07
Control variable: Media	3.59E+03	1.35E+03	2.66	0.008	0.00
Control variable: Homes	1.43E+03	8.45E+02	1.7	0.09	0.00
Control variable: Presentations	3.14E+05	1.06E+05	2.96	0.003	0.00
Control variable: Brochures	5.26E+02	3.08E+02	1.71	0.087	0.00
Control variable: Hazard	6.64E+05	3.02E+05	2.2	0.028	0.00
Media: current month	-4.12E+03	1.34E+03	-3.08	0.002	-0.17
Homes: current month	-1.29E+03	8.41E+02	-1.53	0.125	-0.03
Presentations: current month	-2.97E+05	1.06E+05	-2.82	0.005	-0.23
Brochures: current month	-6.62E+02	3.03E+02	-2.18	0.029	-0.14
Hazard: current month	-6.35E+05	3.00E+05	-2.11	0.035	-0.12
Intercept	-2.97E+01	1.46E+01	-2.03	0.042	
Log Likelihood	-890.5587				
$Prob > chi^2$	0.0000				
Psuedo-R ²	0.7193				

	Regional	Statewide
	Allocations	Allocations
	Cost + Loss	Cost + Loss
	(\$ Million)	(\$ Million)
Current (Base Case)	325	325
Change Prevention Spending Alone,	318	318
Proportional Allocation	516	510
Change RxFire Amount Alone,		
Proportional Allocation, Price	292	306
Responsive RxFire		
Change Prevention Spending Alone,	210	210
Equal Allocation	318	318
Change RxFire Amount Alone, Equal	292	306
Allocation, Price Responsive RxFire	292	300
Change Prevention Spending with Rx		
Fire, Proportional Allocation, Price	287	301
Responsive RxFire		
Change Prevention Spending with Rx		
Fire, Equal Allocation, Price Responsive	287	301
RxFire		
Change Prevention Spending Alone,	202	
Equal Allocation, No Budget Change	323	
Change Prevention Spending Alone,		
Proportional Allocation, No Budget	324	
Change		

598 Table 5. Cost plus loss totals under alternative assumptions and state variables.

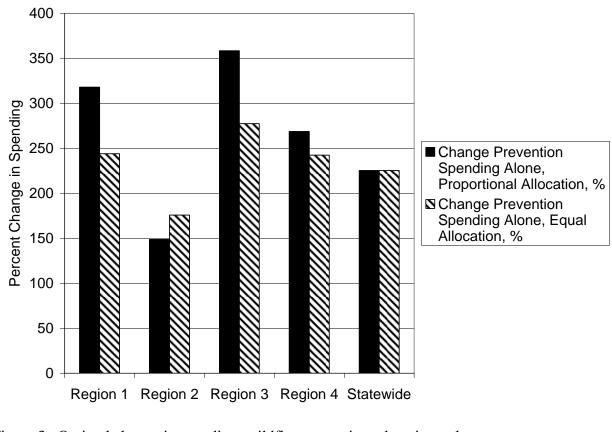




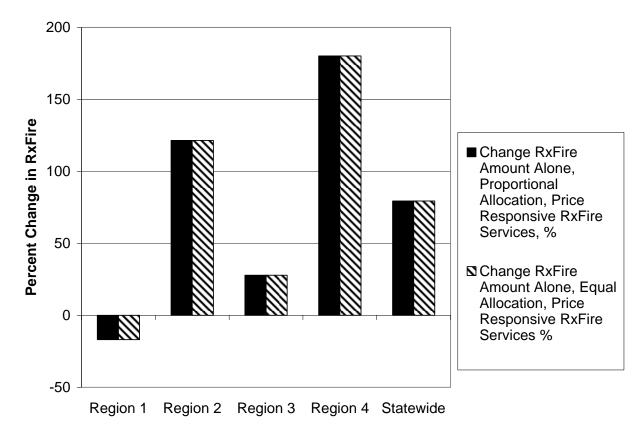


608 Figure 2. Trends of percent deviation from average monthly count of media pubic service

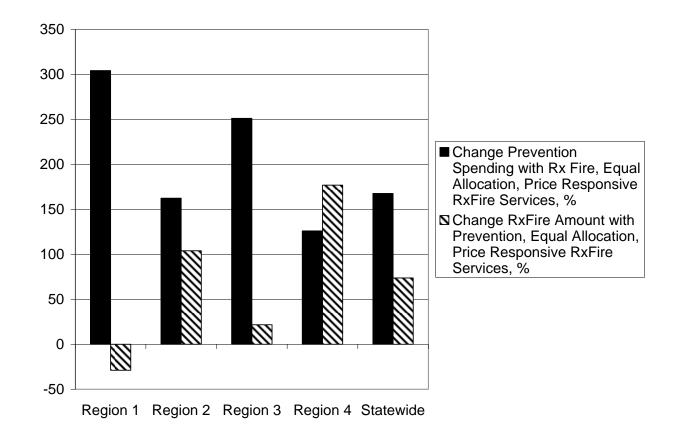
- announcements (PSAs); homes visited; presentations offered; fliers, brochures, and CDs
- 610 distributed; WUI hazard assessments given, prescribed fire fuel treatments (for hazard
- 611 reduction); and targeted unintentional ignitions.



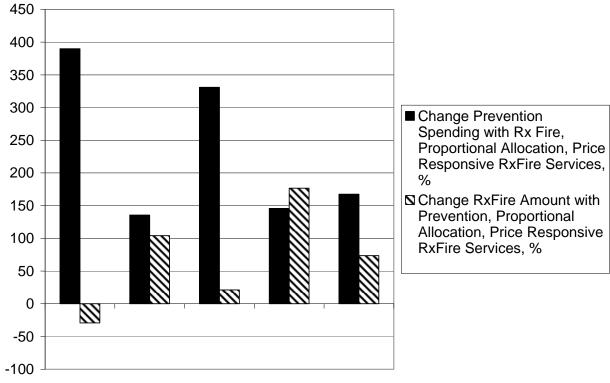
614 Figure 3. Optimal change in spending: wildfire prevention education only.



618 Figure 4. Optimal change in area treated: prescribed fire fuel treatments only.



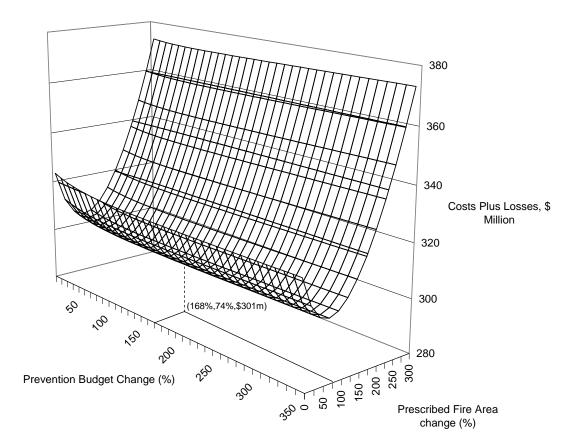
- 622 Figure 5. Optimal change in wildfire mitigation effort: wildfire prevention education (assuming
- 623 equal allocation across regions) and prescribed fire fuel treatments.



Region 1 Region 2 Region 3 Region 4 Statewide

626 Figure 6. Optimal change in wildfire mitigation effort: wildfire prevention education (assuming

- 627 proportional allocation across regions) and prescribed fire fuel treatments.



632 Figure 7. Fire mitigation tradeoff: wildfire prevention education versus prescribed fire fuel

633 treatments.